STA220H1: The Practice of Statistics I

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Please turn on your videos :)



Figure 1: [picture source]

Announcements

- 1. Midterm 2 is next week at 7:20-9:40 PM in EX 100.
- 2. Same rules. Bring your ID.
- 3. Online review session at 6:00-7:00PM, you can stay in EX 100.

Agenda for today

- ▶ Recap: CLT, confidence intervals
- ▶ Statistical testing: H_0 and H_a , process, p-value

We want to study the average life expectancy in Canada μ . We take a sample of *n* people, record their ages of death

 x_1,\ldots,x_n

and compute the sample mean \bar{x} .

We claim that it is an **estimate** of the average life expectancy in Canada.

 $\bar{x} \approx \mu$

How confident are we in our estimate? Can we use our sample to find a range for μ ?

Recap: central limit theorem

Central limit theorem: for *n* large enough

$$ar{X}$$
 approximately \sim Normal $\left(\mu, rac{\sigma^2}{n}
ight)$

When CLT is true?

X₁,..., X_n should be independent and identically distributed
 If X₁,..., X_n are normal then X̄ is exactly normal
 If n > 30 then X̄ is approximately normal

CLT:

$$ar{X}$$
 approximately \sim Normal $\left(\mu, rac{\sigma^2}{n}
ight)$

Standardization:

$$rac{ar{X}-\mu}{\sigma/\sqrt{n}}$$
 approximately \sim *Normal* $(0,1)$

Distribution table:

$$P\left(-1.96 \le rac{ar{X}-\mu}{\sigma/\sqrt{n}} \le 1.96
ight) = 0.95$$

With probability 0.95, the population parameter μ belongs to

$$\left[\bar{X} - 1.96 \cdot \frac{\sigma}{\sqrt{n}}, \bar{X} + 1.96 \cdot \frac{\sigma}{\sqrt{n}}\right]$$

How to use the sample to construct the confidence interval? If σ is known

$$x_1,\ldots,x_n\Rightarrow \bar{x}$$

and the 95% confidence interval is

$$\left[ar{x} - 1.96 \cdot rac{\sigma}{\sqrt{n}}, ar{x} + 1.96 \cdot rac{\sigma}{\sqrt{n}}
ight]$$

How to use the sample to construct the confidence interval? If σ is unknown we estimate it $s \approx \sigma$

$$x_1,\ldots,x_n\Rightarrow \bar{x},s$$

and could probably use the 95% confidence interval

$$\left[ar{x} - 1.96 \cdot rac{s}{\sqrt{n}}, ar{x} + 1.96 \cdot rac{s}{\sqrt{n}}
ight]$$

However it is not accurate...

Both $\bar{X} = \frac{1}{n} \sum_{i=1}^{n} X_i$ and $S^2 = \frac{1}{n-1} \sum_{i=1}^{n} (X_i - \bar{X})^2$ are RVs! Standardization:

$$rac{ar{X}-\mu}{S/\sqrt{n}}$$
approximately $\sim t_{n-1}$

Distribution table:

$$P\left(-...\leq \frac{\bar{X}-\mu}{S/\sqrt{n}}\leq ...\right)=0.95$$

With probability 0.95, the population parameter μ belongs to

$$\left[\bar{X} - \dots \cdot \frac{S}{\sqrt{n}}, \bar{X} + \dots \cdot \frac{S}{\sqrt{n}}\right]$$

How to use the sample to construct the confidence interval? If σ is unknown we estimate it $s \approx \sigma$

$$x_1,\ldots,x_n\Rightarrow \bar{x},s$$

and the 95% confidence interval is

$$\left[\bar{x} - \dots \cdot \frac{s}{\sqrt{n}}, \bar{x} + \dots \cdot \frac{s}{\sqrt{n}}\right]$$

where ... is found from the distribution table.

Recap: t-distribution

Normal: ... = 1.96 t with df = 3: ... = 3.18 t with df = 10: ... = 2.23 t with df = 50: ... = 2.01



Confidence intervals

Known
$$\sigma$$
, 95% CI is $\left[\bar{x} - 1.96 \cdot \frac{\sigma}{\sqrt{n}}, \bar{x} + 1.96 \cdot \frac{\sigma}{\sqrt{n}}\right]$
Unknown σ , 95% CI is $\left[\bar{x} - \dots \cdot \frac{s}{\sqrt{n}}, \bar{x} + \dots \cdot \frac{s}{\sqrt{n}}\right]$

Properties:

- CI is centered at \bar{x}
- CI covers μ with high probability
- Cl size decreases with the growth of n
- CI size decreases with the decrease in confidence
- CI bounds depend on the sample $x_1, ..., x_n$
- If σ is know the CI does not depend on $x_1, ..., x_n$

Exercise

Is σ known for these confidence intervals?

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70	80	90	100					
age of death (years)								

Confidence intervals: interpretation

95% confidence means that for 95% samples CI will cover μ .



Confidence intervals: interpretation

Alternative view:

- We have *n* random variables X_1, \ldots, X_n
- $\bar{X} = \frac{1}{n} \sum_{i=1}^{n} X_i$ and $S^2 = \frac{1}{n-1} \sum_{i=1}^{n} (X_i \bar{X})^2$ are random variables
- ► CI bounds $LB = \bar{X} ... \cdot \frac{S}{\sqrt{n}}$ and $UB = , \bar{X} + ... \cdot \frac{S}{\sqrt{n}}$ are random variables
- ► Each CI is a realization of [LB, UB]

95% confidence means that [LB,UB] will cover μ with probability 0.95.

Exercise

We want to study the average height of people in Canada. We took 500 samples of size 50 and used each sample to compute 90% confidence interval (500 CIs in total). How many of them do you think will not contain the true average height?

Statistical tests use data to answer questions about the population.

Caffeine causes a dramatic increase in the heart rate!

For 30 participants of the experiment, the heart rates before and after coffee was measured the difference was computed. The average value for the difference is 5 bpm.

Why can't we just compare the average difference to zero?

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Goal: determine whether the collected data provides enough evidence for us to believe in a claim about the theoretical world.

Step 1: state your null hypothesis and the alternative hypothesis.

$$\blacktriangleright \text{ Null } H_0: \mu = \mu_0$$

- One sided alternative $H_a: \mu > \mu_0$ or $H_a: \mu < \mu_0$
- Two sided alternative $H_a: \mu \neq \mu_0$

Do our data provide enough evidence against the null?

Step 1: state your null hypothesis and the alternative hypothesis.

 H_0 : the before and after coffee heart rates are the same H_a : the after coffee heart rates is higher than the before one

What are μ and μ_0 here?

Step 2: summarize the data into a test statistic.

• Test statistic is constructed assuming that H_0 is true

How extreme is our test statistic assuming that H_0 is true?

Step 2: summarize the data into a test statistic.

For 30 participants of the experiment, the heart rates before and after coffee was measured the difference was computed. The sample mean is 5 bpm, the sample standard deviation is 0.5.

$$t_{obs} = \frac{\bar{x} - \mu}{s/\sqrt{n}}$$

tobs = (5-0)/(0.5*sqrt(30)) tobs

[1] 1.825742

Step 3: compute p-value.

- It is a number between 0 and 1
- It quantifies how strong is the evidence against H_0
- The smaller the value the stronger the evidence that our data contradict H₀

P-value measures how likely the observed data would be IF the null hypothesis is true.

Step 3: compute **p-value**. Recall that $T = \frac{\bar{X} - \mu}{S/\sqrt{n}} \sim t_{n-1}$ Thus we can find *p-value* = $P(T > t_{obs})$ pt(tobs, df = 29, lower.tail = FALSE)

[1] 0.03910166

Step 4: draw the conclusion.

- If p-value is small we have enough evidence against H_0
- We reject H_0 in favor of H_a
- We say that the result is statistically significant

How small should be p-value?

> The smaller the more significant evidence we have to reject H_0

General rule: pre-select some significance level α and check if $\textit{p-value} < \alpha$

Statisticians prefer *p*-value < 0.05

Step 4: draw the conclusion.

p-value < 0.05, thus we can reject H_0 in favor of H_a .

The after coffee heart rates is higher than the before one!

Statistical testing and tails

- The student was randomly guessing on the exam!
- A student took a test with 100 Yes/No questions. They received the tests results and they got 65 questions correctly.

How to compute p-value?

Step 1: state your **null** hypothesis and the **alternative** hypothesis. $H_0: p = 0.5$ and $H_a: p \neq 0.5$

Step 2: summarize the data into a test statistic.

Since
$$Z = \frac{\bar{X}-p}{\sqrt{p(1-p)/n}} \sim Normal(0,1)$$
 we use

$$z_{obs} = rac{0.65 - 0.5}{\sqrt{0.5(1 - 0.5)/100}}$$

zobs = (0.65-0.5)/sqrt(0.5 * (1-0.5)/100)
zobs

[1] 3

Step 3: compute p-value.

The probability to observe such an extreme statistic is *p*-value $= P(|Z| > |z_{obs}|)$

2*pnorm(zobs, lower.tail = FALSE)

```
## [1] 0.002699796
```

Step 4: draw the conclusion.

p-value < 0.05, thus the student did not guess on the exam!

Statistical testing and tails

Lottery is scamming people!

The lottery company claims that 10% of their tickets win. A customer bought 500 tickets and won only 30 times.

How the procedure will change?

Step 1: state your **null** hypothesis and the **alternative** hypothesis. $H_0: p = 0.1$ and $H_a: p < 0.1$

Step 2: summarize the data into a test statistic.

$$z_{obs} = \frac{0.06 - 0.1}{\sqrt{0.1(1 - 0.1)/500}}$$

zobs = (0.06-0.1)/sqrt(0.1 * (1-0.1)/500) zobs

[1] -2.981424

Step 3: compute p-value.

The probability to observe such an extreme statistic is *p*-value $= P(Z < z_{obs})$

pnorm(zobs, lower.tail = TRUE)

[1] 0.001434556

Step 4: draw the conclusion.

p-value < 0.05, thus the probability to win is less than announced!

Statistical testing and tails

What if p-value is more than 0.05?

- We do not have enough evidence to reject H_0
- This is not the same as to accept H_0 !

TO DO

- 1. Module 8. The Process of Statistical Tests
- 2. Quiz 8 due Monday (March 13) @ 11:59 PM (EST)
- 3. Practice Problem Set 8